Insurance Benefits

Lecture: Weeks 6-8



An introduction

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- Central theme: to quantify the value today of a (random) amount to be paid at a random time in the future.
 - main application is in life insurance contracts, but could be applied in other contexts, e.g. warranty contracts.
- Generally computed in two steps:
 - **(**) take the present value (PV) random variable, $b_T v_T$; and
 - **②** calculate the expected value $E[b_T v_T]$ for the average value this value is referred to as the Actuarial Present Value (APV).
- In general, we want to understand the entire distribution of the PV random variable $b_T v_T$:
 - it could be highly skewed, in which case, there is danger to use expectation.
 - other ways of summarizing the distribution such as variances and percentiles/quantiles may be useful.

A simple illustration

Consider the simple illustration of valuing a three-year term insurance policy issued to age 35 where if he dies within the first year, a \$1,000 benefit is payable at the end of his year of death.

If he dies within the second year, a \$2,000 benefit is payable at the end of his year of death. If he dies within the third year, a \$5,000 benefit is payable at the end of his year of death.

Assume a constant discount rate of 5% and the following extract from a mortality table:

x	q_x
35	0.005
36	0.006
37	0.007
38	0.008

Calculate the APV of the benefits.



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Chapter summary

- Life insurance
 - benefits payable contingent upon death; payment made to a designated beneficiary
 - actuarial present values (APV)
 - actuarial symbols and notation
- Insurances payable at the moment of death
 - continuous
 - level benefits, varying benefits (e.g. increasing, decreasing)
- Insurances payable at the end of year of death
 - discrete
 - level benefits, varying benefits (e.g. increasing, decreasing)
- Chapter 4 (Dickson, et al.) both 1st/2nd ed.

The present value random variable

- Denote by Z, the present value random variable.
- This gives the value, at policy issue, of the benefit payment. Issue age is usually denoted by x.
- In the case where the benefit is payable at the moment of death, Z clearly depends on the time-until-death T. For simplicity, we drop the subscript x for age-at-issue.
- It is $Z = b_T v_T$ where:
 - b_T is called the benefit payment function
 - v_T is the discount function
- In the case where we have a constant (fixed) interest rate, then $v_T = v^T = (1+i)^{-T} = e^{-\delta T}$.

Fixed term life insurance

- An *n*-year term life insurance provides payment if the insured dies within *n* years from issue.
- For a unit of benefit payment, we have

$$b_T = egin{cases} 1, & T \leq n \ 0, & T > n \end{cases}$$
 and $v_T = v^T.$

• The present value random variable is therefore

$$Z = \begin{cases} v^T, & T \le n \\ 0, & T > n \end{cases} = v^T I(T \le n)$$

where $I(\cdot)$ is called indicator function. E[Z] is called the APV of the insurance.

• Actuarial notation:

$$\bar{A}^1_{x:\overline{n}|} = \mathsf{E}[Z] = \int_0^n v^t f_x(t) dt = \int_0^n v^t{}_t p_x \mu_{x+t} dt.$$

Rule of moments

• The j-th moment of the distribution of Z can be expressed as:

$$\mathsf{E}[Z^{j}] = \int_{0}^{n} v^{tj} {}_{t} p_{x} \mu_{x+t} dt = \int_{0}^{n} e^{-(j\delta)t} {}_{t} p_{x} \mu_{x+t} dt.$$

- This is actually equal to the APV but evaluated at the force of interest $j\delta$.
- In general, we have the following rule of moment:

$$\mathsf{E}[Z^j]_{@ \delta_t} = \mathsf{E}[Z]_{@ j\delta_t}.$$

• For example, the variance can be expressed as

$$\operatorname{Var}[Z] = {}^2\bar{A}^1_{x:\overline{n}} - \left(\bar{A}^1_{x:\overline{n}}\right)^2.$$



Whole life insurance

- For a whole life insurance, benefits are payable following death at any time in the future.
- Here, we have $b_T = 1$ so that the present value random variable is $Z = v^T$.
- APV notation for whole life: $\bar{A}_x = \mathsf{E}[Z] = \int_0^\infty v^t {}_t p_x \mu_{x+t} dt.$
- Variance (using rule of moments):

$$\mathsf{Var}[Z] = {}^2 ar{A}_x - ig(ar{A}_xig)^2$$
 .

- Whole life insurance is the limiting case of term life insurance as $n \to \infty$.
- Note also that if the benefit amount is not 1, but say $b_T = b$, then $\mathsf{E}[Z] = b\bar{A}_x$ and that $\mathsf{Var}[Z] = b^2 \left[{}^2\bar{A}_x \left(\bar{A}_x\right)^2 \right]$.

Pure endowment insurance

- For an *n*-year pure endowment insurance, a benefit is payable at the end of *n* years if the insured survives at least *n* years from issue.
- Here, we have $b_T = \begin{cases} 0, & T \le n \\ 1, & T > n \end{cases}$ and $v_T = v^n$ so that the PV r.v. is $\begin{pmatrix} 0 & T \le n \\ 0 & T \le n \end{cases}$

$$Z = \begin{cases} 0, & T \le n \\ v^n, & T > n \end{cases}.$$

• APV for pure endowment: $A_{x:\overline{n}|} = {}_{n}E_{x} = v^{n}{}_{n}p_{x}.$

• Variance (using rule of moments):

$$\mathsf{Var}[Z] = v^{2n}_{n} p_x \cdot {}_n q_x = {}^2A_{x:\overline{n}|} - \left(A_{x:\overline{n}|}\right)^2.$$

• Sometimes, we can also express the present value random variable based on an indicator function:

$$Z = v^n I(T_x > n),$$

where I(E) is 1 if the event E is true, and 0 otherwise.

Endowment insurance

• For an *n*-year endowment insurance, a benefit is payable if death is within *n* years or if the insured survives at least *n* years from issue, whichever occurs first.

• Here, we have
$$b_T = 1$$
 and $v_T = \begin{cases} v^T, & T \le n \\ v^n, & T > n \end{cases}$ so that the PV r.v. is
$$Z = \begin{cases} v^T, & T \le n \\ v^n, & T > n \end{cases}$$

- It is easy to see that we can re-write Z as $Z = v^{\min(T,n)}$.
- APV endowment: $\bar{A}_{x:\overline{n}|} = \bar{A}_{x:\overline{n}|}^1 + A_{x:\overline{n}|}$
- Variance (using rule of moments):

$$\operatorname{Var}[Z] = {}^2 \bar{A}_{x:\overline{n}|} - \left(\bar{A}_{x:\overline{n}|}\right)^2.$$



Deferred insurance

• For an *n*-year deferred whole insurance, a benefit is payable if the insured dies at least *n* years following issue.

• Here, we have $b_T = \begin{cases} 0, & T \leq n \\ 1, & T > n \end{cases}$ and $v_T = v^T$ so that the PV r.v. is $Z = \begin{cases} 0, & T \leq n \\ v^T, & T > n \end{cases}$.

- APV for deferred insurance: ${}_{n|}\bar{A}_{x} = \int_{n}^{\infty} v^{t}{}_{t}p_{x}\mu_{x+t}dt.$
- Variance (using rule of moments):

$$\mathsf{Var}[Z] = {}_{n|}^2 \bar{A}_x - \left({}_{n|} \bar{A}_x\right)^2.$$

Constant force of mortality - all throughout life

Assume mortality is based on a constant force, say μ , and interest is also based on a constant force of interest, say δ .

- Find expressions for the APV for the following types of insurances:
 - whole life insurance;
 - n-year term life insurance;
 - n-year endowment insurance; and
 - *m*-year deferred life insurance.
- Check out the (corresponding) variances for each of these types of insurance.

[Details in class]



De Moivre's law

Find expressions for the APV for the same types of insurances in the case where you have:

• De Moivre's law.



For a whole life insurance of \$1,000 on (x) with benefits payable at the moment of death, you are given:

$$\delta_t = \begin{cases} 0.04, & 0 < t \le 10\\ 0.05, & t > 10 \end{cases}$$

and

$$\mu_{x+t} = \begin{cases} 0.006, & 0 < t \le 10 \\ 0.007, & t > 10 \end{cases}$$

Calculate the actuarial present value for this insurance.



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Equivalent probability calculations

We can also compute probabilities of Z as follows. Consider the present value random variable Z for a whole life issued to age x. For $0 < \alpha < 1$, the following is straightforward:

$$\begin{split} \Pr[Z \leq \alpha] &= & \Pr[e^{-\delta T_x} \leq \alpha = \Pr[-\delta T_x \leq \log(\alpha)] \\ &= & \Pr[T_x > -(1/\delta)\log(\alpha)] = {}_u p_x, \end{split}$$

where

$$u = (1/\delta)\log(1/\alpha) = \log(1/\alpha)^{1/\delta}.$$

- Consider the case where $\alpha = 0.75$ and $\delta = 0.05$. Then $u = \log(1/0.75)^{1/0.05} = 5.753641$.
- Thus, the probability $\Pr[Z \le 0.75]$ is equivalent to the probability that (x) will survive for another 5.753641 years.

Insurances with varying benefits

Туре	b_T	Ζ	APV
Increasing whole life	$\lfloor T+1 \rfloor$	$\lfloor T+1 \rfloor v^T$	$\left(I\bar{A}\right)_x$
Whole life increasing <i>m</i> -thly	$\lfloor Tm+1\rfloor /m$	$v^T \left\lfloor Tm + 1 \right\rfloor / m$	$\left(I^{(m)}\bar{A}\right)_x$
Constant increasing whole life	Т	Tv^T	$\left(\bar{I}\bar{A} ight)_x$
Decreasing <i>n</i> -year term	$\begin{cases} n - \lfloor T \rfloor, & T \le n \\ 0, & T > n \end{cases}$	$\begin{cases} (n - \lfloor T \rfloor) v^T, & T \le n \\ 0, & T > n \end{cases}$	$\left(D\bar{A}\right)^1_{x:\overline{n}}$

* These items will be discussed in class.



For a whole life insurance on (50) with death benefits payable at the moment of death, you are given:

- Mortality follows De Moivre's law with $\omega=110.$
- $b_t = 10000(1.10)^t$, for $t \ge 0$
- $\delta = 5\%$
- $\bullet~Z$ denotes the present value random variable for this insurance.

Calculate E[Z] and Var[Z].

Can you find an explicit expression for the distribution function of Z, i.e. $\Pr[Z \leq z]?$



Insurances payable at EOY of death

- For insurances payable at the end of the year (EOY) of death, the PV r.v. Z clearly depends on the curtate future lifetime K_x .
- It is $Z = b_{K+1}v_{K+1}$.
- To illustrate, consider an *n*-year term insurance which pays benefit at the end of year of death:

$$b_{K+1} = egin{cases} 1, & K = 0, 1, \dots, n-1 \ 0, & ext{otherwise} \end{cases}$$
 , $v_{K+1} = v^{K+1}$,

and therefore

$$Z = \begin{cases} v^{K+1}, & K = 0, 1, \dots, n-1 \\ 0, & \text{otherwise} \end{cases}$$

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• APV of *n*-year term:

$$A_{x:\overline{n}|}^{1} = \mathsf{E}[Z] = \sum_{k=0}^{n-1} v^{k+1}{}_{k|}q_{x} = \sum_{k=0}^{n-1} v^{k+1}{}_{k}p_{x} \cdot q_{x+k}$$

• Rule of moments also apply in discrete situations. For example,

$$\mathsf{Var}[Z] = {}^2A^1_{x:\overline{n}|} - \left(A^1_{x:\overline{n}|}\right)^2,$$

where

$${}^{2}A^{1}_{x:\overline{n}|} = \mathsf{E}\big[Z^{2}\big] = \sum_{k=0}^{n-1} e^{-2\delta(k+1)} {}_{k} p_{x} \cdot q_{x+k}.$$

(Discrete) whole life insurance

Consider a whole life insurance which pays benefit at the end of year of death (for life):

$$b_{K+1} = 1$$
, $v_{K+1} = v^{K+1}$, and $Z = v^{K+1}$.

• APV:
$$A_x = \mathsf{E}[Z] = \sum_{k=0}^{\infty} v^{k+1}{}_k q_x = \sum_{k=0}^{\infty} v^{k+1}{}_k q_x \cdot q_{x+k}$$

• Applying rule of moments,

$$\operatorname{Var}[Z] = {}^{2}A_{x} - (A_{x})^{2} \,,$$

where

$$^{2}A_{x}=\mathsf{E}\left[Z^{2}\right] =\sum_{k=0}^{\infty}e^{-2\delta\left(k+1\right) }{}_{k}p_{x}\cdot q_{x+k}. \label{eq:alpha}$$

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(Discrete) endowment life insurance

• The APV of a (discrete) endowment life insurance is the sum of the APV of a (discrete) term and a pure endowment:

$$A_{x:\overline{n}|} = A_{x:\overline{n}|}^1 + A_{x:\overline{n}|}$$

• The policy pays a death benefit of \$1 at the end of the year of death, if death is prior to the end of n years, and a benefit of \$1 if the insured survives at least n years.

• In effect, we have
$$b_{K+1} = 1$$
 and $v_{K+1} = \begin{cases} v^{K+1}, & K \le n-1 \\ v^n, & K \ge n \end{cases}$ so that the PV r.v. is $Z = \begin{cases} v^{K+1}, & K \le n-1 \\ v^n, & K \ge n \end{cases}$.

• Here $Z = v^{\min(K+1,n)}$ and one can also apply the rule of moments to evaluate the corresponding variance.



Recursive relationships

- The following will be derived/discussed in class:
 - whole life insurance: $A_x = vq_x + vp_x A_{x+1}$
 - term insurance: $A_{x:\overline{n}|}^1 = vq_x + vp_x A_{x+1:\overline{n-1}|}^1$
 - endowment insurance: $A_{x:\overline{n}|} = vq_x + vp_x A_{x+1:\overline{n-1}|}$



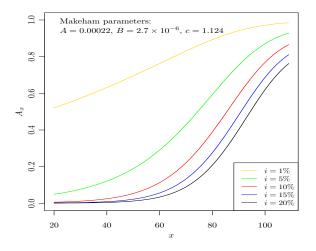


Figure : Actuarial Present Value of a discrete whole life insurance for various interest rate assumptions



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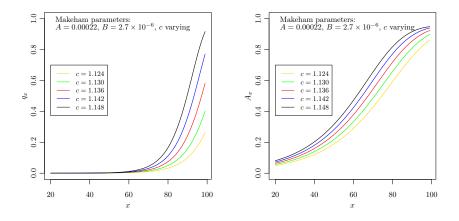


Figure : Actuarial Present Value of a discrete whole life insurance for various mortality rate assumptions with interest rate fixed at 5%

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For a whole life insurance of 1 on (41) with death benefit payable at the end of the year of death, let Z be the present value random variable for this insurance.

You are given:

- i = 0.05;
- $p_{40} = 0.9972;$
- $A_{41} A_{40} = 0.00822$; and
- ${}^{2}A_{41} {}^{2}A_{40} = 0.00433.$

Calculate Var[Z].



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Other forms of insurance

- Deferred insurances
- Varying benefit insurances
- Very similar to the continuous cases
- You are expected to read and understand these other forms of insurances.
- It is also useful to understand the various (possible) recursion relations resulting from these various forms.



Illustration of varying benefits

For a special life insurance issued to (45), you are given:

- Death benefits are payable at the end of the year of death.
- The benefit amount is \$100,000 in the in the first 10 years of death, decreasing to \$50,000 after that until reaching age 65.
- An endowment benefit of \$100,000 is paid if the insured reaches age 65.
- There are no benefits to be paid past the age of 65.
- Mortality follows the Illustrative Life Table at i = 6%.

Calculate the actuarial present value (APV) for this insurance.



For a whole life insurance issued to age 40, you are given:

- Death benefits are payable at the moment of death.
- The benefit amount is \$1,000 in the first year of death, increasing by \$500 each year thereafter for the next 3 years, and then becomes level at \$5,000 thereafter.
- Mortality follows the Illustrative Life Table at i = 6%.
- Deaths are uniformly distributed over each year of age.

Calculate the APV for this insurance.

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Insurance payable m-thly

Insurances payable *m*-thly

- Consider the case where we have just one-year term and the benefit is payable at the end of the *m*-th of the year of death.
- We thus have

$$A_{1}^{(m)}_{x:\overline{1}} = \sum_{r=0}^{m-1} v^{(r+1)/m} \cdot {}_{r/m} p_{x} \cdot {}_{1/m} q_{x+r/m}.$$

• We can show that under the UDD assumption, this leads us to:

$$A_{1}^{(m)}_{x:\overline{1}} = \frac{i}{i^{(m)}} A_{x:\overline{1}}^{1}.$$

• In general, we can generalize this to:

$$A_{1}^{(m)}_{x:\overline{n}} = \frac{i}{i^{(m)}} A_{x:\overline{n}}^{1}$$

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Other types of insurances with m-thly payments

- For other types, we can also similarly derive the following (with the UDD assumption):
 - whole life insurance: $A_x^{(m)} = \frac{i}{i^{(m)}}A_x$
 - deferred life insurance: ${}_{n|}A_x^{(m)} = rac{i}{i^{(m)}n|}A_x$
 - endowment insurance: $A_{x:\overline{n}|}^{(m)} = \frac{i}{i(m)}A_{x:\overline{n}|}^1 + A_{x:\overline{n}|}$



Relationships

Relationships - continuous and discrete

- For some forms of insurances, we can get explicit relationships under the UDD assumption:
 - whole life insurance: $\bar{A}_x = \frac{i}{\delta}A_x$

• term insurance:
$$\bar{A}^1_{x:\overline{n}|} = \frac{i}{\delta} A^1_{x:\overline{n}|}$$

• increasing term insurance: $(I\bar{A})^1_{x:\overline{n}|} = \frac{i}{\delta}(IA)^1_{x:\overline{n}|}$



For a three-year term insurance of 1000 on [50], you are given:

- Death benefits are payable at the end of the quarter of death.
- Mortality follows a select and ultimate life table with a two-year select period:

[x]	$\ell_{[x]}$	$\ell_{[x]+1}$	ℓ_{x+2}	x+2
50	9706	9687	9661	52
51	9680	9660	9630	53
52	9653	9629	9596	54

• Deaths are uniformly distributed over each year of age.

• i = 5%

Calculate the APV for this insurance.



Each of 100 independent lives purchases a single premium 5-year deferred whole life insurance of 10 payable at the moment of death.

You are given:

- $\mu = 0.004$
- $\delta = 0.006$
- F is the aggregate amount the insurer receives from the 100 lives.
- The 95th percentile of the standard Normal distribution is 1.645.

Using a Normal approximation, calculate F such that the probability the insurer has sufficient funds to pay all claims is 0.95.

Suppose interest rate i=6% and mortality is based on the following life table:

	90										
ℓ_x	800	740	680	620	560	500	440	380	320	100	0

Calculate the following:

(a) A_{94} (b) $A_{90:\overline{5}|}^{1}$ (c) $_{3|}A_{92}^{(4)}$, assuming UDD between integral ages (d) $A_{95:\overline{3}|}$



A five-year term insurance policy is issued to (45) with benefit amount of \$10,000 payable at the end of the year of death.

Mortality is based on the following select and ultimate life table:

x	$\ell_{[x]}$	$\ell_{[x]+1}$	$\ell_{[x]+2}$	ℓ_{x+3}	x+3
45	5282	5105	4856	4600	48
46	4753	4524	4322	4109	49
47	4242	4111	3948	3750	50
48	3816	3628	3480	3233	51

Calculate the APV for this insurance if i = 5%.



Other terminologies

Other terminologies and notations used

Expression	Other terms/symbols used
Actuarial Present Value (APV)	Expected Present Value (EPV) Net Single Premium (NSP) single benefit premium
basis	assumptions
interest rate (i)	interest per year effective discount rate
benefit amount (b)	sum insured $\left(S ight)$ death benefit
Expected value of \boldsymbol{Z}	E(Z)
Variance of Z	$\operatorname{Var}(Z) = V[Z]$



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