# Insurance Benefits 

Lecture: Weeks 6-8

## An introduction

- Central theme: to quantify the value today of a (random) amount to be paid at a random time in the future.
- main application is in life insurance contracts, but could be applied in other contexts, e.g. warranty contracts.
- Generally computed in two steps:
(1) take the present value (PV) random variable, $b_{T} v_{T}$; and
(2) calculate the expected value $\mathrm{E}\left[b_{T} v_{T}\right]$ for the average value - this value is referred to as the Actuarial Present Value (APV).
- In general, we want to understand the entire distribution of the PV random variable $b_{T} v_{T}$ :
- it could be highly skewed, in which case, there is danger to use expectation.
- other ways of summarizing the distribution such as variances and percentiles/quantiles may be useful.


## A simple illustration

Consider the simple illustration of valuing a three-year term insurance policy issued to age 35 where if he dies within the first year, a $\$ 1,000$ benefit is payable at the end of his year of death.
If he dies within the second year, a $\$ 2,000$ benefit is payable at the end of his year of death. If he dies within the third year, a $\$ 5,000$ benefit is payable at the end of his year of death.
Assume a constant discount rate of $5 \%$ and the following extract from a mortality table:

| $x$ | $q_{x}$ |
| :---: | :---: |
| 35 | 0.005 |
| 36 | 0.006 |
| 37 | 0.007 |
| 38 | 0.008 |

Calculate the APV of the benefits.

## Chapter summary

- Life insurance
- benefits payable contingent upon death; payment made to a designated beneficiary
- actuarial present values (APV)
- actuarial symbols and notation
- Insurances payable at the moment of death
- continuous
- level benefits, varying benefits (e.g. increasing, decreasing)
- Insurances payable at the end of year of death
- discrete
- level benefits, varying benefits (e.g. increasing, decreasing)
- Chapter 4 (Dickson, et al.) - both 1st/2nd ed.


## The present value random variable

- Denote by $Z$, the present value random variable.
- This gives the value, at policy issue, of the benefit payment. Issue age is usually denoted by $x$.
- In the case where the benefit is payable at the moment of death, $Z$ clearly depends on the time-until-death $T$. For simplicity, we drop the subscript $x$ for age-at-issue.
- It is $Z=b_{T} v_{T}$ where:
- $b_{T}$ is called the benefit payment function
- $v_{T}$ is the discount function
- In the case where we have a constant (fixed) interest rate, then $v_{T}=v^{T}=(1+i)^{-T}=e^{-\delta T}$.


## Fixed term life insurance

- An $n$-year term life insurance provides payment if the insured dies within $n$ years from issue.
- For a unit of benefit payment, we have

$$
b_{T}=\left\{\begin{array}{ll}
1, & T \leq n \\
0, & T>n
\end{array} \text { and } v_{T}=v^{T}\right.
$$

- The present value random variable is therefore

$$
Z=\left\{\begin{array}{ll}
v^{T}, & T \leq n \\
0, & T>n
\end{array}=v^{T} I(T \leq n)\right.
$$

where $I(\cdot)$ is called indicator function. $\mathrm{E}[Z]$ is called the APV of the insurance.

- Actuarial notation:

$$
\bar{A}_{x: \bar{n} \mid}^{1}=\mathrm{E}[Z]=\int_{0}^{n} v^{t} f_{x}(t) d t=\int_{0}^{n} v^{t}{ }_{t} p_{x} \mu_{x+t} d t
$$

## Rule of moments

- The $j$-th moment of the distribution of $Z$ can be expressed as:

$$
\mathrm{E}\left[Z^{j}\right]=\int_{0}^{n} v^{t j}{ }_{t} p_{x} \mu_{x+t} d t=\int_{0}^{n} e^{-(j \delta) t}{ }_{t} p_{x} \mu_{x+t} d t
$$

- This is actually equal to the APV but evaluated at the force of interest $j \delta$.
- In general, we have the following rule of moment:

$$
\mathrm{E}\left[Z^{j}\right] @ \delta_{t}=\mathrm{E}[Z]_{@ j \delta_{t}} .
$$

- For example, the variance can be expressed as

$$
\operatorname{Var}[Z]={ }^{2} \bar{A}_{x: \bar{n} \mid}^{1}-\left(\bar{A}_{x: \bar{n}}^{1}\right)^{2}
$$

## Whole life insurance

- For a whole life insurance, benefits are payable following death at any time in the future.
- Here, we have $b_{T}=1$ so that the present value random variable is $Z=v^{T}$.
- APV notation for whole life: $\bar{A}_{x}=\mathrm{E}[Z]=\int_{0}^{\infty} v^{t}{ }_{t} p_{x} \mu_{x+t} d t$.
- Variance (using rule of moments):

$$
\operatorname{Var}[Z]={ }^{2} \bar{A}_{x}-\left(\bar{A}_{x}\right)^{2}
$$

- Whole life insurance is the limiting case of term life insurance as $n \rightarrow \infty$.
- Note also that if the benefit amount is not 1 , but say $b_{T}=b$, then $\mathrm{E}[Z]=b \bar{A}_{x}$ and that $\operatorname{Var}[Z]=b^{2}\left[{ }^{2} \bar{A}_{x}-\left(\bar{A}_{x}\right)^{2}\right]$.


## Pure endowment insurance

- For an $n$-year pure endowment insurance, a benefit is payable at the end of $n$ years if the insured survives at least $n$ years from issue.
- Here, we have $b_{T}=\left\{\begin{array}{ll}0, & T \leq n \\ 1, & T>n\end{array}\right.$ and $v_{T}=v^{n}$ so that the PV r.v. is

$$
Z= \begin{cases}0, & T \leq n \\ v^{n}, & T>n\end{cases}
$$

- APV for pure endowment: $A_{x: \frac{1}{n}}={ }_{n} E_{x}=v^{n}{ }_{n} p_{x}$.
- Variance (using rule of moments):

$$
\operatorname{Var}[Z]=v^{2 n}{ }_{n} p_{x} \cdot{ }_{n} q_{x}={ }^{2} A_{x: \frac{1}{n \mid}}-\left(A_{x: \frac{1}{n}}\right)^{2} .
$$

- Sometimes, we can also express the present value random variable based on an indicator function:

$$
Z=v^{n} I\left(T_{x}>n\right),
$$

where $I(E)$ is 1 if the event $E$ is true, and 0 otherwise.

## Endowment insurance

- For an $n$-year endowment insurance, a benefit is payable if death is within $n$ years or if the insured survives at least $n$ years from issue, whichever occurs first.
- Here, we have $b_{T}=1$ and $v_{T}=\left\{\begin{array}{ll}v^{T}, & T \leq n \\ v^{n}, & T>n\end{array}\right.$ so that the PV r.v. is $Z=\left\{\begin{array}{ll}v^{T}, & T \leq n \\ v^{n}, & T>n\end{array}\right.$.
- It is easy to see that we can re-write $Z$ as $Z=v^{\min (T, n)}$.
- APV endowment: $\bar{A}_{x: \bar{n}}=\bar{A}_{x: \bar{n}}^{1}+A_{x: \frac{1}{n}}$.
- Variance (using rule of moments):

$$
\operatorname{Var}[Z]={ }^{2} \bar{A}_{x: \bar{n} \mid}-\left(\bar{A}_{x: \bar{n})}\right)^{2} .
$$

## Deferred insurance

- For an $n$-year deferred whole insurance, a benefit is payable if the insured dies at least $n$ years following issue.
- Here, we have $b_{T}=\left\{\begin{array}{ll}0, & T \leq n \\ 1, & T>n\end{array}\right.$ and $v_{T}=v^{T}$ so that the PV r.v. is

$$
Z= \begin{cases}0, & T \leq n \\ v^{T}, & T>n\end{cases}
$$

- APV for deferred insurance: ${ }_{n \mid} \bar{A}_{x}=\int_{n}^{\infty} v^{t}{ }_{t} p_{x} \mu_{x+t} d t$.
- Variance (using rule of moments):

$$
\operatorname{Var}[Z]={ }_{n \mid}^{2} \bar{A}_{x}-\left({ }_{n \mid} \bar{A}_{x}\right)^{2}
$$

## Constant force of mortality - all throughout life

Assume mortality is based on a constant force, say $\mu$, and interest is also based on a constant force of interest, say $\delta$.

- Find expressions for the APV for the following types of insurances:
- whole life insurance;
- $n$-year term life insurance;
- $n$-year endowment insurance; and
- $m$-year deferred life insurance.
- Check out the (corresponding) variances for each of these types of insurance.
[Details in class]


## De Moivre's law

Find expressions for the APV for the same types of insurances in the case where you have:

- De Moivre's law.


## Illustrative example 1

For a whole life insurance of $\$ 1,000$ on $(x)$ with benefits payable at the moment of death, you are given:

$$
\delta_{t}= \begin{cases}0.04, & 0<t \leq 10 \\ 0.05, & t>10\end{cases}
$$

and

$$
\mu_{x+t}= \begin{cases}0.006, & 0<t \leq 10 \\ 0.007, & t>10\end{cases}
$$

Calculate the actuarial present value for this insurance.

## Equivalent probability calculations

We can also compute probabilities of $Z$ as follows. Consider the present value random variable $Z$ for a whole life issued to age $x$. For $0<\alpha<1$, the following is straightforward:

$$
\begin{aligned}
\operatorname{Pr}[Z \leq \alpha] & =\operatorname{Pr}\left[e^{-\delta T_{x}} \leq \alpha=\operatorname{Pr}\left[-\delta T_{x} \leq \log (\alpha)\right]\right. \\
& =\operatorname{Pr}\left[T_{x}>-(1 / \delta) \log (\alpha)\right]={ }_{u} p_{x},
\end{aligned}
$$

where

$$
u=(1 / \delta) \log (1 / \alpha)=\log (1 / \alpha)^{1 / \delta}
$$

- Consider the case where $\alpha=0.75$ and $\delta=0.05$. Then $u=\log (1 / 0.75)^{1 / 0.05}=5.753641$.
- Thus, the probability $\operatorname{Pr}[Z \leq 0.75]$ is equivalent to the probability that $(x)$ will survive for another 5.753641 years.


## Insurances with varying benefits

| Type | $b_{T}$ | $Z$ | APV |
| :--- | :--- | :--- | :--- |
| Increasing <br> whole life | $\lfloor T+1\rfloor$ | $\lfloor T+1\rfloor v^{T}$ | $(I \bar{A})_{x}$ |
| Whole life <br> increasing $m$-thly | $\lfloor T m+1\rfloor / m$ | $v^{T}\lfloor T m+1\rfloor / m$ | $\left(I^{(m)} \bar{A}\right)_{x}$ |
| Constant increasing <br> whole life | $T$ | $T v^{T}$ | $(\bar{I} \bar{A})_{x}$ |
| Decreasing <br> $n$-year term | $\left\{\begin{array}{llll}n-\lfloor T\rfloor, & T \leq n \\ 0, & T>n\end{array}\right.$ | $\begin{cases}(n-\lfloor T\rfloor) v^{T}, & T \leq n \\ 0, & T>n\end{cases}$ | $(D \bar{A})_{x: \bar{n}}^{1}$ |

* These items will be discussed in class.


## Illustrative example 2

For a whole life insurance on (50) with death benefits payable at the moment of death, you are given:

- Mortality follows De Moivre's law with $\omega=110$.
- $b_{t}=10000(1.10)^{t}$, for $t \geq 0$
- $\delta=5 \%$
- $Z$ denotes the present value random variable for this insurance.

Calculate $\mathrm{E}[Z]$ and $\operatorname{Var}[Z]$.
Can you find an explicit expression for the distribution function of $Z$, i.e. $\operatorname{Pr}[Z \leq z]$ ?

## Insurances payable at EOY of death

- For insurances payable at the end of the year (EOY) of death, the PV r.v. $Z$ clearly depends on the curtate future lifetime $K_{x}$.
- It is $Z=b_{K+1} v_{K+1}$.
- To illustrate, consider an $n$-year term insurance which pays benefit at the end of year of death:

$$
b_{K+1}=\left\{\begin{array}{ll}
1, & K=0,1, \ldots, n-1 \\
0, & \text { otherwise }
\end{array}, v_{K+1}=v^{K+1}\right.
$$

and therefore

$$
Z= \begin{cases}v^{K+1}, & K=0,1, \ldots, n-1 \\ 0, & \text { otherwise }\end{cases}
$$

## - continued

- APV of $n$-year term:

$$
A_{x: \bar{n} \mid}^{1}=\mathrm{E}[Z]=\sum_{k=0}^{n-1} v^{k+1}{ }_{k \mid} q_{x}=\sum_{k=0}^{n-1} v^{k+1}{ }_{k} p_{x} \cdot q_{x+k}
$$

- Rule of moments also apply in discrete situations. For example,

$$
\operatorname{Var}[Z]={ }^{2} A_{x: \bar{n} \mid}^{1}-\left(A_{x: \bar{n}}^{1}\right)^{2},
$$

where

$$
{ }^{2} A_{x: \bar{n} \mid}^{1}=\mathrm{E}\left[Z^{2}\right]=\sum_{k=0}^{n-1} e^{-2 \delta(k+1)}{ }_{k} p_{x} \cdot q_{x+k} .
$$

## (Discrete) whole life insurance

Consider a whole life insurance which pays benefit at the end of year of death (for life):

$$
b_{K+1}=1, v_{K+1}=v^{K+1}, \text { and } Z=v^{K+1}
$$

- APV: $A_{x}=\mathrm{E}[Z]=\sum_{k=0}^{\infty} v^{k+1}{ }_{k \mid} q_{x}=\sum_{k=0}^{\infty} v^{k+1}{ }_{k} q_{x} \cdot q_{x+k}$
- Applying rule of moments,

$$
\operatorname{Var}[Z]={ }^{2} A_{x}-\left(A_{x}\right)^{2},
$$

where

$$
{ }^{2} A_{x}=\mathrm{E}\left[Z^{2}\right]=\sum_{k=0}^{\infty} e^{-2 \delta(k+1)}{ }_{k} p_{x} \cdot q_{x+k}
$$

## (Discrete) endowment life insurance

- The APV of a (discrete) endowment life insurance is the sum of the APV of a (discrete) term and a pure endowment:

$$
A_{x: \bar{n} \mid}=A_{x: \bar{n} \mid}^{1}+A_{x: \left.\frac{1}{n} \right\rvert\,}
$$

- The policy pays a death benefit of $\$ 1$ at the end of the year of death, if death is prior to the end of $n$ years, and a benefit of $\$ 1$ if the insured survives at least $n$ years.
- In effect, we have $b_{K+1}=1$ and $v_{K+1}=\left\{\begin{array}{ll}v^{K+1}, & K \leq n-1 \\ v^{n}, & K \geq n\end{array}\right.$ so that the PV r.v. is $Z=\left\{\begin{array}{ll}v^{K+1}, & K \leq n-1 \\ v^{n}, & K \geq n\end{array}\right.$.
- Here $Z=v^{\min (K+1, n)}$ and one can also apply the rule of moments to evaluate the corresponding variance.


## Recursive relationships

- The following will be derived/discussed in class:
- whole life insurance: $A_{x}=v q_{x}+v p_{x} A_{x+1}$
- term insurance: $A_{x: \bar{n} \mid}^{1}=v q_{x}+v p_{x} A_{x+1: \overline{n-1}}^{1}$
- endowment insurance: $A_{x: \bar{n}}=v q_{x}+v p_{x} A_{x+1: \overline{n-1}}$


Figure : Actuarial Present Value of a discrete whole life insurance for various interest rate assumptions


Figure : Actuarial Present Value of a discrete whole life insurance for various mortality rate assumptions with interest rate fixed at 5\%

## Illustrative example 3

For a whole life insurance of 1 on (41) with death benefit payable at the end of the year of death, let $Z$ be the present value random variable for this insurance.
You are given:

- $i=0.05$;
- $p_{40}=0.9972$;
- $A_{41}-A_{40}=0.00822$; and
- ${ }^{2} A_{41}-{ }^{2} A_{40}=0.00433$.

Calculate $\operatorname{Var}[Z]$.

## Other forms of insurance

- Deferred insurances
- Varying benefit insurances
- Very similar to the continuous cases
- You are expected to read and understand these other forms of insurances.
- It is also useful to understand the various (possible) recursion relations resulting from these various forms.


## Illustration of varying benefits

For a special life insurance issued to (45), you are given:

- Death benefits are payable at the end of the year of death.
- The benefit amount is $\$ 100,000$ in the in the first 10 years of death, decreasing to \$50,000 after that until reaching age 65.
- An endowment benefit of $\$ 100,000$ is paid if the insured reaches age 65.
- There are no benefits to be paid past the age of 65 .
- Mortality follows the Illustrative Life Table at $i=6 \%$.

Calculate the actuarial present value (APV) for this insurance.

## Illustrative example 4

For a whole life insurance issued to age 40, you are given:

- Death benefits are payable at the moment of death.
- The benefit amount is $\$ 1,000$ in the first year of death, increasing by $\$ 500$ each year thereafter for the next 3 years, and then becomes level at $\$ 5,000$ thereafter.
- Mortality follows the Illustrative Life Table at $i=6 \%$.
- Deaths are uniformly distributed over each year of age.

Calculate the APV for this insurance.

## Insurances payable $m$-thly

- Consider the case where we have just one-year term and the benefit is payable at the end of the $m$-th of the year of death.
- We thus have

$$
\underset{x: 1\rceil}{A_{1}^{(m)}}=\sum_{r=0}^{m-1} v^{(r+1) / m} \cdot{ }_{r / m} p_{x} \cdot{ }_{1 / m} q_{x+r / m} .
$$

- We can show that under the UDD assumption, this leads us to:

$$
\underset{\substack{x: 1]}}{A_{i}^{(m)}}=\frac{i}{i^{(m)}} A_{x: 1 \mid}^{1} .
$$

- In general, we can generalize this to:

$$
\underset{\substack{x: \bar{n}}}{A_{1}^{(m)}}=\frac{i}{i^{(m)}} A_{x: \bar{n} \mid}^{1} .
$$

## Other types of insurances with $m$-thly payments

- For other types, we can also similarly derive the following (with the UDD assumption):
- whole life insurance: $A_{x}^{(m)}=\frac{i}{i^{(m)}} A_{x}$
- deferred life insurance: ${ }_{n \mid} A_{x}^{(m)}=\frac{i}{i^{(m)}}{ }_{n \mid} A_{x}$
- endowment insurance: $A_{x: \overline{n \mid}}^{(m)}=\frac{i}{i^{(m)}} A_{x: \bar{n}}^{1}+A_{x: \bar{n}}$


## Relationships - continuous and discrete

- For some forms of insurances, we can get explicit relationships under the UDD assumption:
- whole life insurance: $\bar{A}_{x}=\frac{i}{\delta} A_{x}$
- term insurance: $\bar{A}_{x: \bar{n} \mid}^{1}=\frac{i}{\delta} A_{x: \bar{n} \mid}^{1}$
- increasing term insurance: $(I \bar{A})_{x: \bar{n}}^{1}=\frac{i}{\delta}(I A)_{x: \bar{n}}^{1}$


## Illustrative example 5

For a three-year term insurance of 1000 on [50], you are given:

- Death benefits are payable at the end of the quarter of death.
- Mortality follows a select and ultimate life table with a two-year select period:

| $[x]$ | $\ell_{[x]}$ | $\ell_{[x]+1}$ | $\ell_{x+2}$ | $x+2$ |
| :---: | :---: | :---: | :---: | :---: |
| 50 | 9706 | 9687 | 9661 | 52 |
| 51 | 9680 | 9660 | 9630 | 53 |
| 52 | 9653 | 9629 | 9596 | 54 |

- Deaths are uniformly distributed over each year of age.
- $i=5 \%$

Calculate the APV for this insurance.

## Illustrative example 6

Each of 100 independent lives purchases a single premium 5 -year deferred whole life insurance of 10 payable at the moment of death.
You are given:

- $\mu=0.004$
- $\delta=0.006$
- $F$ is the aggregate amount the insurer receives from the 100 lives.
- The 95th percentile of the standard Normal distribution is 1.645.

Using a Normal approximation, calculate $F$ such that the probability the insurer has sufficient funds to pay all claims is 0.95 .

## Illustrative example 7

Suppose interest rate $i=6 \%$ and mortality is based on the following life table:

| $x$ | 90 | 91 | 92 | 93 | 94 | 95 | 96 | 97 | 98 | 99 | 100 |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\ell_{x}$ | 800 | 740 | 680 | 620 | 560 | 500 | 440 | 380 | 320 | 100 | 0 |

Calculate the following:
(a) $A_{94}$
(b) $A_{90: 5}^{1}$
(c) ${ }_{3 \mid} A_{92}^{(4)}$, assuming UDD between integral ages
(d) $A_{95: 31}$

## Illustrative example 8

A five-year term insurance policy is issued to (45) with benefit amount of $\$ 10,000$ payable at the end of the year of death.
Mortality is based on the following select and ultimate life table:

| $x$ | $\ell_{[x]}$ | $\ell_{[x]+1}$ | $\ell_{[x]+2}$ | $\ell_{x+3}$ | $x+3$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 45 | 5282 | 5105 | 4856 | 4600 | 48 |
| 46 | 4753 | 4524 | 4322 | 4109 | 49 |
| 47 | 4242 | 4111 | 3948 | 3750 | 50 |
| 48 | 3816 | 3628 | 3480 | 3233 | 51 |

Calculate the APV for this insurance if $i=5 \%$.

## Other terminologies and notations used

| Expression | Other terms/symbols used |
| :--- | :--- |
| Actuarial Present Value (APV) | Expected Present Value (EPV) <br> Net Single Premium (NSP) <br> single benefit premium |
| basis | assumptions |
| interest rate $(i)$ | interest per year effective <br> discount rate |
| benefit amount $(b)$ | sum insured $(S)$ <br> death benefit |
| Expected value of $Z$ | $\mathrm{E}(Z)$ |
| Variance of $Z$ | $\operatorname{Var}(Z) \quad V[Z]$ |

